

Do Dividends and Earnings Ratios Explain the Excess Equity Premium? A Study of the Indian Stock Market

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Abstract

The excess market return earned by investors over the risk-free rate is termed as equity premium. Mehra and Prescott (1985) found a very high equity premium of 6% for the US market, which was very puzzling for various reasons. Several theories have guided the curiosity of this puzzle attributable to the consumption behaviour pattern and habits formation of individual investors. We used aggregate ratios of firms' dividends and earnings as a proxy for investors' consumption behaviour to estimate the equity premium for a cross-sectional portfolio of firms using in-sample and out-sample data. We found that, though, the dividend-price and earnings-price ratios acted as best in-sample predictors of excess equity returns. The estimated risk-premium varied across the cross-section of firms.

Keywords: Equity Premium, Excess Risk-Premium, Dividend-Price, Earnings-Price, Earnings-Yield, Dividend-Yield, Firm Size, BM ratio

Introduction

The rational expectation framework proposed that the risk-averse representative agents in the general equilibrium always tried to maximise their expected value of the discounted stream of cash-flows generated through an expected utility model (Lucas, 1978). It is criticised as it failed to account for the large return differentials in the expected stock returns over the risk-free rate in the US market documented by Mehra and Prescott (1985). They estimated a large equity premium of 6.18% and the risk-free rate of 0.8% vis-à-vis the historically observed premium of 0.35% for the US market, which was termed as the equity premium puzzle. The rationalists attributed the large equity premium to macro-economic factors viz., the deferred consumption by risk-averse investors (John Y. Campbell & Cochrane, 1999; Weil, 1989) whereas the behaviourists proposed that the behaviour of individual investors viz., the external habit formation (Constantinides, 1990), and the behaviour of representative agent according to the prospect theory (Benartzi & Thaler, 1995) caused anomalous equity premium.

Studies by (Campbell & Cochrane, 2000; Fama & French, 1988; Welch & Goyal, 2008) used the aggregate dividends and earnings of firms as a proxy for market consumption to estimate equity premium. In our study, we used aggregate dividends ratios and earnings ratios viz., dividend-price ratio, dividend-yield ratio, earnings-price ratio and earnings-yield ratio, for estimating the excess equity returns for the

Indian market using the data of the NSE listed firms. The study is different in three things, viz., - 1) we used the best-fit linear model for estimation of equity returns and adjusted for the non-linearity using a best-fit conditional-variance model, 2) portfolio-based estimation of excess equity returns, and 3) we used in-sample and out-sample estimations for the robustness of estimation models. We found that the estimated excess risk premium varied across the cross-section of firms. Both the dividend-price and earnings-price ratios estimated the in-sample equity premium; the out-sample estimations were not better failing to prove the robustness of the estimation models. This paper is structured to discuss the literature review, theoretical framework for the selection of variables and hypothesised relationship between variables in section 2; data and methodology in section 3; discussion of results in section 4; and findings and conclusion in section 5.

Literature Review

Theories on Equity premium puzzle

The equity premium is the extra risk premium earned by investors by making an investment in a portfolio of risky securities over risk-free securities. Mehra and Prescott (1985) found a large equity premium of 6-8% vis-à-vis the historically observed equity premium of 0.35% for the US market, which was puzzling. A few studies tried to explain the equity premium puzzle by proposing new theories. The rationalist theory of Weil (1989) attributed the substantial equity premium to the deferred consumption behaviour of risk-averse investors attempting to generate higher per capita consumption growth rate over the rate offered by the risk-free Treasury bills. He suggested the equity premium puzzle transformed into the risk-free rate puzzle as the risk-averse investors under Lucas (1978) framework deferred their consumption by saving more even when the risk-free rate of returns offered was meagre. Campbell and Cochrane (1999) caused anomalous equity premium.

Studies by (Campbell & Cochrane, 2000; Fama & French, 1988; Welch & Goyal, 2008) used the aggregate explained that the equity premium puzzle through a consumption-based model where the aggregate consumption behaviour of representative agents perfectly correlated with the business cycle. The pro-cyclical consumption declined towards the habit during the business trough, and the counter-cyclical equity premium puzzle increased due to the cross-sectional variations in the wealth distribution of the heterogeneous representative agents.

The behaviourist theory of Constantinides (1990) explained the equity premium puzzle and the risk-free rate puzzle through the model of 'habit-formation' of the representative agents, where risk-averse agents didn't consider the effect the current consumption on future

preferences, but their utility depended on the levels of past consumption. The model failed to explain the consumption behaviour of wealthy investors, pension funds and endowments at the aggregate market level (Benartzi & Thaler, 1995). The duo proposed that investors behaved according to the prospect theory (Kahneman & Tversky, 1979), where, the risk-averse investors demanded a higher premium for assuming a higher variability in the securities returns due to their high sensitivity to losses. Thus, the highly risk-averse investors demand higher risk-premium as compensation for undertaking investments in the highly-risky securities.

Estimation of excess risk-premium

Mehra and Prescott (1985) used the real S&P index return and the estimated real interest rate for the estimation of the equity premium. The security risk used in the estimation model was measured as a covariance of security returns with the per capita consumption, a proxy for consumption stream of investors under the rational utility framework.

Welch (2000) surveyed finance professionals regarding estimation of the equity premium and found that there was neither a proper explanation for high values of equity premium nor any consensus on how to forecast equity premium. Thus there is an overlap of techniques used for estimation of expected stock returns and the excess equity returns, i.e. the equity premium.

The seminal literature on the estimation of stock returns viz., the Capital Asset Pricing Model (CAPM) (Black, 1972; Lintner, 1965; Sharpe, 1964) established a linear relationship between the expected returns on securities and the market risk, where, the market risk is measured by beta in the general equilibrium model. The CAPM proposed that beta alone explained the expected stock returns given that markets are efficient. If the efficient markets hypothesis were to hold, the stocks were to be priced rationally, then the systemic differences in stock returns are attributed to the differences in risk. The market capitalisation of firm (Banz, 1981), the earnings-price (EP) ratio (Basu, 1983; Chan, Hamao, & Lakonishok, 1991) and the book equity to market equity (BM) ratio (Chan, Jegadeesh, & Lakonishok, 1995; Fama & French, 1992, 1993, 1996, 2006, 2012) along with the market beta significantly contributed for explaining the cross-sectional variation of expected stock returns.

Campbell and Cochrane (2000) used the consumption-based model (Campbell & Cochrane, 1999) and the CAPM (Black, 1972; Lintner, 1965; Sharpe, 1964) to explain the time-varying expected returns using the dividend-price ratio as a proxy for market consumption. They found that, though, both the models estimated conditional asset returns, however, the portfolio-based models better

approximated the unconditional asset returns.

Campbell and Cochrane (1999, 2000) observed that changes in dividend explained more than half of the variation in stock returns and, variations in the aggregate dividend-price ratio is due to variations in aggregated expected excess returns. Berk (1995) found a strong correlation between expected stock return and dividend-yield, along with other non-systemic firm variables, failed the CAPM to account for cross-sectional differences in expected stock returns.

Fama and French (2002) suggested that any variable that is co-integrated with stock price can be used to estimate expected stock return, but the ratios should be mean-reverting and stationary. They observed multicollinearity between the dividend-price ratio and earnings-price ratio. Their study found that estimation of expected returns using the firm's fundamentals outperformed any other estimation on the ground of lower standard error and Sharpe ratio.

Dimson, Marsh, and Staunton (2008) decomposed the equity risk premium into three components viz., the level of dividends, the growth in dividends and the effects of stock prices on dividend-price ratio. Fama and French (1988) used dividend yields for the estimation of expected stock returns. Fama and French (2002), Welch and Goyal (2008) used aggregate earnings and dividends ratios (dividend-price, dividend-yield, earnings-price, and earnings-yield ratios) for estimation of excess risk-premium; the theoretical arguments on asset pricing proposed that an average estimated stock return is the sum of the average dividend yield and the average rate of capital gain. These studies have predominantly used linear estimation technique for the expected equity premium. The dividend-yield ratio and the dividend-price ratio were found to be the dominant predictors of future returns using artificial neural networks (Wong, Hassan, & Feroz, 2007; Welch & Goyal, 2008). Welch and Goyal (2008) accounted for the non-linearity in the predictor variables viz., dividends ratios. Claus and Thomas (2001) and Gebhardt, Hvidkjaer, and Swaminathan (2005) used valuation models involving dividends and earnings ratios to estimate unconditional expected returns.

Siegel (1992) and Siegel and Thaler (1997) acknowledged that the standard asset pricing models could not explain the higher equity premium. Damodaran (2009, 2012) noted that the variation in expected equity premium vis-à-vis the actual historical equity premium of about 3% to 12% is due to the choice of different estimation periods, differences in risk-free rates and market indices, and differences in the way returns are averaged over time.

The literature analysis identifies that there is an overlap of estimation techniques used for expected stock returns, and

equity premium, and no proper consensus on the estimation technique for equity premium. Previous studies have predominantly used linear estimation techniques and valuation models for estimation of conditional and unconditional expected returns. It is found that dividends ratios acted as powerful predictors of expected stock returns. However, the time-dependent, non-linear and non-stationary dividends ratios contradict the proposition of stationary and mean-reverting predictor variables resulting in sparse estimation. Thus there is a strong need for studies to explore robust estimation techniques for equity premium. In our study, we perform a linear estimation of equity premium after adjusting for the time-dependent characteristics of the predictor variables by using the best-fit conditional and unconditional estimation models. We test for the robustness of the models by testing their out-sample performance. We use the portfolio-based estimation of equity premium for cross-sectional portfolios formed based on EP ratio, Firm size and the BM ratio rather than using firm-level data.

Data and Methodology

We have used the monthly data of NSE listed firms, i.e. NSE 500 firms in the study. The period considered in the study is between 2004 and 2015. We formed different cross-sectional portfolios for the estimation of the equity premium.

Formation of Cross-sectional Portfolios

The NSE 500 firms have been categorised into different cross-sectional portfolios, sorted based on EP ratio, BM ratio and the market capitalisation of the firms. The annual averages of the ratios were cumulated and ranked in descending order. The top 10 percent and the bottom 10 percent of firms under each category are considered for the analysis. Thus, we have considered six portfolios of 50 stocks each viz., EP High (top 10% EP firms), EP Low (Bottom 50 EP firms), Growth firms (top 10% BM firms), Value firms (Bottom 50 BM firms), MK High (top 50 market cap firms), and MK Low (bottom 10% market cap firms).

Variables of the estimation model

We collected the monthly closing prices of NSE 500 stocks, their EP ratio, market capitalisation and the BM ratio from the CMIE Prowess database. The CCIL 90 days Treasury-bill index is used as a proxy for the risk-free rate. The T-bill index data is downloaded from the CCIL website, the sister website of the NSE. The variables used for regression are explained in Table 1

Table 1: Variables used for estimation of the equity premium

Notation	Formula	Description
Variables		
P_t	$P_t = \frac{1}{N} \sum_{i=1}^N p_{it}$ Where p_{it} = stock price of an i th stock at the time 't', N = No. of stocks in the portfolio	Equal-weighted portfolio price at time 't.'
P_{t-1}	$P_t (-1)$	Lag price of the equal-weighted portfolio at time 't'
D_t	$D_t = \frac{1}{N} \sum_{i=1}^N d_{it}$ Where d_{it} = dividends of the i th stock at the time 't';	Aggregate dividends of all the stock in the portfolio at time 't'
E_t	$E_t = \frac{1}{N} \sum_{i=1}^N eps_{it}$ Where eps_t = earnings per share of the i th stock at the time 't';	Aggregate earnings of all the stocks in the portfolio at time 't'
Rf_t	$Rf_t = \ln\left(\frac{r_t}{r_{t-1}}\right)$ Where, r_t = risk-free rate of risk-free portfolio at time 't'	Risk-free return at time 't' of the portfolio of risk-free treasury bills
Rf_{t-1}	$Rf_t (-1)$	The risk-free rate at time 't-1' of the portfolio of risk-free treasury bills
MP_t	$MP_t = \frac{1}{N} \sum_{i=1}^N p_{i,t} * n_i$ Where p_{it} = price of the i th stock at time t; n = no. of outstanding stocks of the i th stock;	The market capitalisation of the equal-weighted portfolio at time 't'
Dependent Variable		
$EQPM_t$	$EQPM_t = \left[\left(\frac{P_t + D_t}{P_{t-1}} \right) - \left(\frac{Rf_t}{Rf_{t-1}} \right) \right]$	The equity premium is the difference between the return on the

Predictor Variables		portfolio of risky securities (capital gain + dividends) over the portfolio of risk-free treasury bills.
DP_t	$DP_t = \left(\frac{D_t}{MP_t} \right)$	Dividend-price is the ratio of an aggregate dividend of the portfolio to a total market capitalisation of the portfolio at time 't'
DY_t	$DY_t = \left(\frac{D_t}{MP_{t-1}} \right)$	Dividend-yield is measured as the ratio of an aggregate dividend of the portfolio at the time 't' to the total market capitalisation of the portfolio at the time 't-1'
EP_t	$EP_t = \left(\frac{E_t}{MP_t} \right)$	Earnings-price is the ratio of aggregate earnings of the portfolio to the total market capitalisation of the portfolio at the time 't'
EY_t	$EY_t = \left(\frac{E_t}{MP_{t-1}} \right)$	Earnings-yield is measured as the ratio of aggregate earnings of the portfolio at the time 't' to the total market capitalisation of the portfolio at time 't-1'

□

Estimation models

We have used Wong, Hassan, and Feroz (2007) and Welch and Goyal (2008) method for the estimation of equity premium using the dividends ratios and earnings ratios as predictor variables. We carried out both the in-sample and out-sample estimations for the robustness of the model by dividing the dataset two viz., the in-sample (70%, i.e. 2004 to 2013 data) and out-of-sample data (30%, i.e. 2013 – 2015 data). We used OLS multiple regression techniques for mean estimation and the ARMA-GARCH estimation for conditional mean-variance estimations, for taking care

of serial correlation and heteroscedasticity problems in the residuals. The base model is represented mathematically in (1). The ARMA (P, Q) -GARCH (p,q) representations are given in (2) and (3).

$$EQPM_t = \alpha_1 + \beta_1 DP_t + \beta_2 DY_t + \beta_3 EP_t + \beta_4 EY_t + \varepsilon_t \quad (1)$$

Where, $\alpha_1, \beta_1, \beta_2, \beta_3$, and β_4 are estimates of OLS multiple regression, ε_t is error of AR (p) order

$$EQPM_t = \alpha_1 + \beta_1 DP_t + \beta_2 DY_t + \beta_3 EP_t + \beta_4 EY_t + \delta_1 AR(1) + \dots + \delta_p AR(p) + \theta_1 MA(1) + \dots + \theta_q MA(q) \quad (2)$$

$$h_t^2 = \alpha_0 + \gamma_1 \varepsilon_{t-1}^2 + \dots + \gamma_q \varepsilon_{t-q+1}^2 + \tau_1 h_{t-1} + \dots + \tau_p h_{t-p+1} \quad (3)$$

Where, $\delta_1, \dots, \delta_p$ represent coefficients of AR (P); $\theta_1, \dots, \theta_q$ represent coefficients of MA (Q); $\gamma_1, \dots, \gamma_q$ represent coefficients of ARCH (q) and $\tau_1, \dots, \tau_p$ represent coefficients of GARCH (p)

In-Sample estimation

In the in-sample estimation, we have done the conditional mean forecasting of equity premium using the multiple OLS regression with best-fit ARMA-GARCH model for adjusting for the non-linearity in residuals viz., serial correlation and the heteroscedasticity problems. The OLS regression estimates provide the degree of the linear relationship of the equity premium with the predictor variables. The estimates are unbiased if the residuals of the model satisfy the properties of CLRM (Classical Linear Regression Model), i.e. BLUE (Best Linear Unbiased Estimators). The presence of significant autocorrelation and heteroscedasticity in the residuals violate the assumption of i.i.d. (identical independent distribution).

We also checked the series for non-stationarity and multicollinearity issues. The non-stationary series produce spurious regression estimates. In order to test for the non-stationarity of the dependent and independent variables, we used the unit root tests (e-views version 8) viz., the ADF (Augmented Dickey-Fuller) test and the PP (Philip-Peron-Fisher Chi-square) tests, for testing the null hypothesis that the series has a unit root (non-stationary). The null hypothesis is rejected at 5% level of significance if the ADF and PP statistics are higher than the respective critical values.

We performed the VIF test (Variance Inflation Factor) for testing the multicollinearity of the predictor variables in the general estimation model (1) across each portfolio. If $VIF > 10$, then there is no multicollinearity among independent variables. We took EQPM as the dependent variable, and the computed EP, EY, DP and DY ratios are taken as independent variables for carrying out the OLS multiple-regression. The residual diagnostics tests of the residuals of the OLS regression showed auto-correlation and heteroscedasticity problems. We fitted the best-fit ARMA-GARCH model, which was selected based on the SIC (Schwarz Information Criteria) for the conditional mean-variance estimation of the equity premium.

Out-Sample estimation

The out-of-sample forecasting is done using the n-step ahead conditional forecasting of the equity premium using the in-sample estimates. The robustness of the best-fit in-sample model is ascertained by computing RMSE (Root Mean Squared Error), MAE (Mean Absolute Error) and MAPE (Mean Absolute Percentage Error) values for both in-sample and out-sample data. The RMSE, MAE and MAPE predict the forecast accuracy of the estimation model. They are computed using (4), (5) and (6)

$$RMSE = \sqrt{\sum_{t=T+1}^{T+h} \frac{(\hat{y}_t - y_t)^2}{h}} \quad (4)$$

Where h is the sample size

$$MAE = \sum_{t=T+1}^{T+h} \frac{|\hat{y}_t - y_t|}{h} \quad (5)$$

$$MAPE = \sum_{t=T+1}^{T+h} \frac{|\hat{y}_t - y_t|}{h} * 100 \quad (6)$$

The forecast accuracy measures for both in-sample and out-sample data are compared for robustness. A robust estimation model should produce accurate estimates for the out-sample data.

Results and Discussion

Table 2 provides the summary statistics of the variables for cross-sectional portfolios. The aggregate earnings and dividends ratios viz., EP, EY, DP and EY for all the portfolios are found to be non-normally distributed (skewness $\neq 0$ and kurtosis $\neq 3$). The estimated average

equity premium is ranging from 0.7% – 0.8% for the portfolios. The estimated average equity premium is more for BM Low (0.81%) and MK Low (0.80%) and less for BM High (0.71%) and MK High (0.71%) portfolios, implying higher risk-premium for small size (MK Low) and value portfolios (BM Low). The findings support the size effect and value effect (Banz, 1981; Basu, 1983; Chan, Hamao, & Lakonishok, 1991; Chan, Jegadeesh, & Lakonishok, 1995; Fama & French, 1992, 1993, 1996, 2006, 2012)

Table 2: Summary Statistics of Variables

Portfolio	Mean	Median	Std. Dev	Variance	Max	Min	Kurtosis	Skewness	Count
BM High									
EQPM	0.7176	0.7614	0.2520	0.0635	1.1841	0.0000	-0.7712	-0.4032	125
EP	0.1112	0.1203	0.0495	0.0025	0.2685	0.0335	-0.0348	0.4302	125
EY	0.1103	0.1193	0.0499	0.0025	0.2579	0.0000	-0.0566	0.3181	125
DP	0.0315	0.0282	0.0223	0.0005	0.0867	0.0000	-0.9551	0.3972	125
DY	0.0314	0.0279	0.0225	0.0005	0.0867	0.0000	-0.9707	0.4069	125
BM Low									
EQPM	0.8140	0.8566	0.1843	0.0340	1.1660	0.0000	1.6972	-0.7869	125
EP	0.2400	0.2337	0.0769	0.0059	0.4538	0.0986	0.4561	0.6612	125
EY	0.2391	0.2328	0.0780	0.0061	0.4538	0.0000	0.6339	0.4022	125
DP	0.0358	0.0180	0.0330	0.0011	0.1319	0.0001	0.1212	1.1940	125
DY	0.0352	0.0184	0.0328	0.0011	0.1320	0.0000	0.3794	1.2863	125
EP High									
EQPM	0.7321	0.7750	0.2047	0.0419	1.1337	0.0000	-0.1610	-0.4093	125
EP	0.0741	0.0707	0.0303	0.0009	0.1850	0.0205	2.0352	1.2518	125
EY	0.0737	0.0684	0.0308	0.0010	0.1828	0.0000	1.9682	1.1199	125
DP	0.0213	0.0229	0.0111	0.0001	0.0469	0.0009	-1.1079	0.1201	125
DY	0.0213	0.0226	0.0112	0.0001	0.0454	0.0000	-1.1256	0.1154	125
EP Low									
EQPM	0.7927	0.8661	0.2246	0.0504	1.1542	0.0000	-0.1112	-0.7065	125
EP	0.1492	0.1566	0.0924	0.0085	0.4976	0.0255	1.4396	0.8249	125
EY	0.1497	0.1560	0.0936	0.0088	0.4724	0.0000	0.3419	0.6390	125
DP	0.0296	0.0147	0.0363	0.0013	0.1440	0.0001	2.2605	1.8538	125
DY	0.0294	0.0148	0.0367	0.0013	0.1505	0.0000	2.6869	1.9485	125

MK High									
EQPM	0.7178	0.7627	0.1749	0.0306	1.0441	0.0000	0.9063	-0.6048	125
EP	0.1746	0.1552	0.0640	0.0041	0.3752	0.0725	0.5154	0.9683	125
EY	0.1739	0.1583	0.0650	0.0042	0.3410	0.0000	0.1770	0.7329	125
DP	0.0199	0.0123	0.0146	0.0002	0.0708	0.0020	1.7769	1.4060	125
DY	0.0197	0.0129	0.0144	0.0002	0.0705	0.0000	1.7834	1.4021	125
MK Low									
EQPM	0.8003	0.8481	0.2824	0.0798	1.2852	0.0000	-0.8257	-0.3260	125
EP	0.1053	0.1075	0.0681	0.0046	0.2526	0.0083	-1.0650	0.2991	125
EY	0.1052	0.0979	0.0693	0.0048	0.2530	0.0000	-1.0563	0.3059	125
DP	0.0597	0.0287	0.0672	0.0045	0.2746	0.0000	1.1758	1.5231	125
DY	0.0599	0.0294	0.0685	0.0047	0.2891	0.0000	1.5910	1.6081	125

Source: Authors' computation

Table 3 provides the results for the stationarity of the in-sample data. It is observed that the p-values of the ADF test and the PP-Fisher test are significant at 5% level of

significance, i.e. $p < 0.05$, the null-hypothesis of unit-root (non-stationary) is rejected. Therefore, the series is stationary at level.

Table 3: Tests for Stationarity for In-sample data

Method	Statistic	Prob.**	Cross-sections	Observations
Null: Unit root (assumes common unit root process)				
Levin, Lin & Chu t^*	0.40258	0.6564	30	2916
Null: Unit root (assumes individual unit root process)				
Im, Pesaran and Shin W-stat	-1.79004	0.0367	30	2916
ADF - Fisher Chi-square	85.7878	0.0161	30	2916
PP - Fisher Chi-square	156.913	0.0000	30	2940

Source: e-views output

Table 4 describes the results for multi-collinearity of the predictor variables. It is observed that the VIF $>> 10$ for all the variables of the estimation model for all the portfolio.

Therefore, there is no problem with multicollinearity among predictor variables.

Table 4: Test for multicollinearity

OLS Model/Portfolio	Variables	VIF
BM High	EP_BMHIGH	60.20005
	EY_BMHIGH	57.73873
	DP_BMHIGH	102.9117
	DY_BMHIGH	103.3876
BM Low	DP_BMLOW	65.50051
	DY_BMLOW	66.38482
	EP_BMLOW	53.64268
	EY_BMLOW	50.38351
EP High	DP_EPHIGH	93.5656
	DY_EPHIGH	93.44293
	EP_EPHIGH	71.22695
	EY_EPHIGH	67.53738
EP Low	DP_EPLOW	89.83399
	DY_EPLOW	88.97493
	EP_EPLOW	88.21873
	EY_EPLOW	83.54424
MK High	DP_MKHIGH	102.3738
	DY_MKHIGH	104.9051
	EP_MKHIGH	45.94443
	EY_MKHIGH	45.95151
MK Low	DP_MKLOW	17.96127
	DY_MKLOW	16.95436
	EP_MKLOW	98.53914
	EY_MKLOW	95.96268

Source: Authors' computation

Table 5 shows the estimates of the OLS multiple-regression model for the in-sample data. The best-fit ARMA-GARCH estimation technique is used for conditional mean-variance estimation of the equity premium. We found, AR (2) for BMHigh and MKHigh portfolios; AR(1)-GARCH(0,1) for BMLow and EPHigh portfolios; and AR (1) for EPLow and MKLow portfolios, respectively, as the best-fit conditional mean and conditional mean-variance equations. The adjusted R² were 0.97 (BMHigh), 0.81 (BMLow), 0.82 (EPHigh), 0.92 (EPLow), 0.89 (MKHigh) and 0.95 (MKLow). The adjusted R² of 0.97 implies that 97% of the variation in the equity premium is explained by the predictor variables (estimation model) for BMHigh portfolios. The residual diagnostic tests showed no heteroscedasticity and no ARCH effects. The DW statistics

showed serial correlations in the residuals, as $DW \neq 2$ for all portfolios. The F-statistic is significant for all estimation models.

The results showed that all the estimated coefficients are significant at 5% level of significance for BMHigh (except intercept), EPHigh, EPLow, MKHigh and MKLow portfolios. However, for BMLow portfolio, the variation in dividend-price and dividend-yield ratios didn't significantly influence the change in equity premium as the coefficient estimates were not significant at 5% level. It is observed that price-ratios are negatively associated with equity premium, and yield-ratios are positively associated with equity premium for all portfolios (except MKLow, where, DP is positively associated with equity premium).

This implies that a decrease in dividends and earnings at a time 't' reduces the equity premium available for shareholders. However, equity premium increases as the dividend-yield and earnings-yield ratios increase for all the stocks.

Table 5: Estimates of the best-fit multiple regression model (In-sample data)

Dependent variable: EQPM

Best-fit OLS		Coefficient		t-stat/ z-	p-	Adj.				
Model	Variable	t	SE	stat	value	R ²	F-stat	DW		
BM HIGH										
AR(2)	DP	-7.622	1.33	-5.702	0.000	0.975	2	1.906		
			7							
	DY	8.317	1.23	6.743	0.000					
			3							
	EP	-1.995	0.33	-5.960	0.000					
			5							
	EY	1.985	0.26	7.556	0.000					
			3							
AR(1)	C	1.408	1.43	0.979	0.330					
			8							
	AR(1)	0.527	0.07	7.573	0.000					
			0							
	AR(2)	0.462	0.06	6.684	0.000					
			9							
	BM Low									
	GARCH (0,1)	DP	-1.580	1.35	-1.168		0.243	0.805		2.675
3										
DY		2.031	1.78	1.138	0.255					
			4							
EP		-1.336	0.14	-9.536	0.000					
			0							
EY		1.501	0.17	8.561	0.000					
			5							
RESID(-1)^2	C	0.820	0.15	5.447	0.000					
			0							
	AR(1)	0.953	0.04	21.565	0.000					
			4							
			0.22							
	1)^2	0.602	3	2.700	0.007					

EP High								
AR(1) -								
GARCH (0,1)	DP	-10.563	1.556	-6.786	0.000	0.820		2.348
	DY	11.973	1.326	9.032	0.000			
	EP	-2.751	0.452	-6.084	0.000			
	EY	2.495	0.509	4.907	0.000			
	C	0.860	0.177	4.849	0.000			
	AR(1)	0.974	0.020	49.202	0.000			
	RESID(-1)^2	0.959	0.287	3.344	0.001			
EP Low								
AR (1)	DP	-2.588	1.144	-2.261	0.026	0.928	251.116	2.375
	DY	3.574	1.072	3.336	0.001			
	EP	-2.454	0.273	-8.995	0.000			
	EY	2.983	0.287	10.394	0.000			
	C	0.702	0.062	11.358	0.000			
	AR(1)	0.867	0.040	21.785	0.000			
MK High								
AR (2)	DP	-14.370	2.604	-5.517	0.000	0.898	141.518	2.267
	DY	15.566	2.611	5.962	0.000			
	EP	-0.885	0.180	-4.925	0.000			
	EY	1.043	0.182	5.732	0.000			
	C	0.730	0.092	7.946	0.000			
	AR(1)	0.483	0.082	5.863	0.000			
	AR(2)	0.445	0.079	5.656	0.000			
MK Low								
AR (1)	DP	0.544	0.257	2.113	0.037	0.955	417.138	2.605
	DY	0.680	0.257	2.652	0.009			
	EP	-3.625	0.524	-6.920	0.000			
	EY	4.847	0.458	10.593	0.000			
	C	0.601	0.076	7.896	0.000			
	AR(1)	0.871	0.046	19.043	0.000			

Source: e-views output

Notes: Column 1 shows the type of best-fit ARMA-GARCH model used. Column 2 represents the estimated variables. Column 3 shows the coefficients of the estimation. Column 4 shows the standard error of the estimates. Column 5 shows the t-statistic/ z-statistic value of the estimates. The tests are done for the statistical significance of the estimates. Column 6 shows the p-value of the estimates. The p-value < 0.05 implies that the estimates are highly significant at 5% level. Column 7 shows the adjusted R², which shows the goodness-of-fit of the model. Column 8 shows the F-statistic value of the regression, which is a proxy for the significance of the regression (higher F-statistic). Column 9 shows the Durbin-Watson value for serial-correlation in the residuals. DW =2 implies there is no auto-correlation in residuals.

Robustness of the models (Out-sample estimation)

Table 6 shows the forecast accuracy measures for in-sample and out-sample estimations. The predictor variables performed well in in-sample estimation for BMHigh, BMLow, MKHigh portfolios as the forecast errors measured by RMSE, MAE and MAPE were less (<< 10%) compared to that of EP High, EPLow and MKLow portfolios, where RMSE and MAPE were greater. (For EPHigh, RMSE= 12%, MAPE = 17%; EPLow, RMSE = 10%, MAPE=14%; MKLow, RMSE = 10%, MAPE=14% respectively). The out-sample measures viz., RMSE and MAE are greater for BMLow, EPLow, MKHigh and MKLow portfolios indicating the poor out-sample performance of the estimation models. However, the dividends and earnings-ratios acted as best in-sample and out-sample predictors for EPHigh portfolio.

Table 6: Comparison of forecast errors of in -sample and out -sample estimation for all portfolios

Portfolio	ARMA-GARCH type	Obs.	In-sample estimation			Out-sample estimation			
			RMSE	MAE	MAPE	Obs.	RMSE	MAE	MAPE
BM High	AR (2)	96	6.575%	5.098%	11.148%	26	42.549%	3.457%	3.464%
	AR(1) -								
BM Low	GARCH (0,1)	98	8.510%	6.261%	9.375%	28	9.831%	7.727%	7.588%
	AR(1) -								
EP High	GARCH (0,1)	98	12.428%	9.538%	17.177%	26	6.001%	5.301%	5.328%
EP Low	AR (1)	98	10.481%	8.987%	14.054%	26	14.890%	13.815%	13.738%
MK High	AR (2)	97	7.137%	6.027%	9.65%	26	10.533%	8.866%	9.727%
MK Low	AR (1)	98	10.054%	8.641%	14.919%	26	16.190%	14.144%	13.678%

Source: Authors' computation

Findings and Conclusion

The standard financial models failed to account for the large return differentials in the expected stock returns over the risk-free rate, leading to the puzzle of equity premium documented by Mehra and Prescott (1985). The proponents of market efficiency attributed the higher risk-premium to the macro-economic consumption of representative agents, measured by aggregate dividends and earnings of firms (Campbell and Cochrane, 2000; Fama and French, 1988; Goyal and Welch, 2008).

In our study, we estimated cross-sectional equity premiums using aggregate dividends and earnings ratios. The equity premium varies accordingly with firm characteristics, risk-free rates and predictor variables (Damodaran, 2009; 2012). Therefore, we used the cross-sectional portfolios formed based on market capitalisation (firm size), BM ratio and EP ratios. The OLS multivariate regression techniques, along with the best-fit ARMA-GARCH model, were used for the estimation of conditional mean-variance equations of the expected equity premiums for in-sample and out-sample data. We found that both dividends-ratios and earnings-ratios acted as best in-sample predictors for all portfolios except the Value (BMLow) portfolio. The dividends ratios didn't significantly estimate the equity premium for BMLow portfolio, and the results are incongruent with the findings of Goyal and Welch (2008) that time-varying dividend-ratios predicted themselves better than predicting equity premium. The poor out-sample estimation shows that, though the aggregate dividends and earnings partly explained the cross-sectional equity premiums, they acted as weaker predictors given their random walk behaviour (Goyal and Welch, 2008). Thus, our estimation techniques don't contribute to providing any solid explanation for the anomalous behaviour of equity premium nor the uncertainty surrounding the predictability of the expected equity premium. Thus, there is a need for further research to explore robust techniques involving behavioural variables along with firm variables reflecting the economic behaviour to explain the behaviour of excess equity returns better.

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